Quantitative Coupling of Human Antithrombin Ⅲ Mutations with Their Clinical Outcomes* 定量联结人抗凝血酶Ⅱ的变异及其临床表现

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Abstract: This study demonstrated how to convert a protein sequence into a numeric datum, and provided a descriptively probabilistic method to analyze the genotype-phenotype relationship of single gene disorder. The amino-acid distribution probability was used to quantify human antithrombin I mutations, then the cross-impact analysis was used to couple the quantified mutations with their clinical outcomes, and finally the Bayesian equation was used to determine the probability that the antithrombin I deficiency was determined under a mutation. The results showed that a person had a chance larger than 90% of being inherited antithrombin II deficiency when a mutation was found in the antithrombin I.

Key words: amino acid, antithrombin I deficiency, single gene disorder, mutation, distribution probability, Bayes' law, cross-impact analysis

摘要:将蛋白质序列转换成数值数据后,建立一种描述性的概率方法分析单基因疾病的基因型与表型之间的 关系。先用氨基酸分布概率定量变异的抗凝血酶 I,然后用交叉影响分析法联结人抗凝血酶 I的变异及其临 床表现,再用贝叶斯公式计算变异条件下抗凝血酶 I缺乏的发病概率。结果显示当某个人被检测到抗凝血酶 I 变异时将有大于 90%的机会患遗传性抗凝血酶 I缺乏。 关键词:氨基酸 遗传性抗凝血酶 I缺乏 单基因疾病 变异 分布概率 贝叶斯法 交叉影响分析法

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The antithrombin \mathbf{II} is a vitamin K-independent hepatocyte-synthesized protease inhibitor, which binds with and blocks the biologic activity of thrombin and other activated coagulation proteins involved in the clotting cascade^[1].

The building of quantitative relationship between mutations in antithrombin **I** and their clinical outcomes is intriguing, not only because 91 mutations have been documented in UniProtKB/ Swiss-Prot entry^[2], but also it would provide the basis for understanding and modeling of genotypephenotype relationship. An important clinical outcome is the inherited antithrombin \blacksquare deficiency, which was first described in a Norwegian family^[3]. Since then, many families with such deficiency have been described, where this disorder had caused severe venous thrombosis in successive generations, deep venous thrombus^[4,5], cerebral venous thrombosis in neonatal^[6,7], young^[8] and adult patients^[9], and so on.

However, it is understandable that a mutation in antithrombin I could either cause a deficiency or have no clinical effects. In this context, a mutation has a yes-no relationship with its outcome, namely, an event with two possibilities, so the relationship between a single mutation with its outcome is quite simple. However, this would not be the case when there are many mutations involved, because each mutation occurred at different position with different substituted and substituting amino acids, which

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would result in a multiple yes-no relationship. Thus a method is needed in order to treat mutations as changed values rather than changed letters, which represent amino acids in different positions of antithrombin \mathbf{I} .

The simplest method would replace each amino acid in human antithrombin II with a value that could be the physicochemical property of amino acid^[10]. However, the physicochemical properties are not sensitive to mutations, thus our group has developed three methods based on random mechanism to transfer a protein sequence into a numeric sequence since $1999^{[11~14]}$. Hence, each mutation could be quantified with different value, and we could build a quantitative relationship between the mutations in the antithrombin II and inherited antithrombin II deficiency, which was designed as the aim of this study.

1 Materials and methods

1.1 Data

The human antithrombin **I** sequence with its 91 mutations was obtained from UniProtKB/Swiss-Prot entry^[2], of which 82 were missense mutations and 9 deletions.

1.2 Amino-acid distribution probability before and after mutation in antithrombin II

The positions of amino acids in antithrombin \mathbf{I} could be associated with probability that was calculated according to $(r!)/(r_1! \times r_2! \times \cdots \times r_n!) \times (r!)/(q_0! \times q_1! \times \cdots \times q_n!) \times n^{-r[15]}$, where r was the number of amino acids, n was the number of partitions, r_n was the number of amino acids in the n-th partition, q_n was the number of partitions with the same number of amino acids, and ! was the factorial function.

For example, there were five histidines in human antithrombin II, locating at positions 33,97,152,351and 401, respectively. Accordingly, we could imagine dividing the antithrombin II into five partitions because there were five histidines. As the antithrombin II was composed of 464 amino acids, each partition would include about 93 amino acids (464/5 = 92.8). Then we counted how many histidines in each partition (column 3, Table 1), where one partition had two histidines, three partitions contained one histidine, and one partition had none, so we had $r_1 = 1, r_2 = 2, r_3 = 0, r_4 = 1, r_5$ = 1, $q_0 = 1, q_1 = 3, q_2 = 1, q_3 = 0, q_4 = 0, q_5 = 0$, and (5!)/(1! ×2! ×0! ×1! ×1!)×[(5!)/(1! ×3! ×1! ×0! ×0! ×0!)]×5⁻⁵=120/(1×2×1×1× 1)×[120/(1×6×1×1×1×1)]×1/3125 = 0.3840. This was the distribution probability for the histidines in antithrombin **I** before mutation.

. If a mutation occurred related to histidine, its distribution probability would be changes. For example, a mutation in the antithrombin I at position 425 substituted arginine for histidine leading to the type I antithrombin I deficiency^[16]. In the mutant, there were six histidines with respect to their distributions in six partitions (column 5, Table 1), where $r_1 = 1, r_2 = 2, r_3 = 0, r_4 = 0, r_5 = 1, r_6 = 2$, $q_0 = 2, q_1 = 2, q_2 = 2, q_3 = 0, q_4 = 0, q_5 = 0, q_6 = 0,$ i.e. $(6!)/(1! \times 2! \times 0! \times 0! \times 1! \times 2!) \times [(6!)/(1! \times 2!) \times (6!)]$ $(\times 2! \times 2! \times 2! \times 0! \times 0! \times 0! \times 0!)$ 0.3472. Thus, this mutation decreased the distribution probability of histidines to 0.3472 from 0.384 (the last row in Table 1).

Table 1 Distributions and probabilities of arginines and histidines before and after R425H mutation in human antithrombin II

Partition/	Before mutation		After mutation	
Probability	Arginines	Histidines	Arginines	Histidines
I	1	1	1	1
I	0	2	0	2
I	2	0	2	0
N	2	1	2	0
v	1	1	1	1
VI	0	-	0	2
VI	0	-	0	-
VIL	2	-	2	-
IX	1	-	1	-
х	0	-	1	-
XI	2	-	1	-
XI	0	-	0	-
ХШ	1	-	1	-
XN	3	-	3	-
X V	0	-	0	-
хи	0	-	0	-
XVI	2	-	2	-
XVII	0	-	1	-
XIX	1	-	0	· _
xx	0	-	2	-
ХXI	3	-	2	-
XXI	2	-	0	-
XXI	0	-	-	-
Probability	0.0222	0.3840	0.0659	0.3472

Also, this mutation was related to arginines, whose distribution probabilities were 0.0222 and 0.0659 before and after mutation.

The overall effect of this mutation on the antithrombin \mathbf{I} was (0.0659-0.0222)+(0.3472-0.384)=0.0069, that was, the mutation increased the distribution probability for this mutant antithrombin \mathbf{I} . In this way, we had the amino-acid distribution probability for each mutation and its documented clinical outcome. Actually we have used this method in many our previous studies related to different subjects^[11-14.17~22].

2 Results and discussion

After above computation, we had the amino-acid distribution probabilities of normal human antithrombin I and its 91 mutants, of which 85 mutations were documented inherited as antithrombin I deficiency. Thus, we could use the cross-impact analysis to couple the mutations with the occurrence/nonoccurrence of inherited antithrombin I deficiency together, because the cross-impact analysis was particularly suited for coupling two relevant events^[20,23~28].

Fig. 1 showed the scheme based on cross-impact analysis for events defined above. At the level of amino-acid distribution probability, P(2) and $P(\overline{2})$ were the decreased and increased probabilities induced by mutations, and 36 and 55 mutations resulted in the distribution probability decreased and increased, respectively. At the level of clinical outcome: (i) $P(1 \mid P(1 \mid$ $\overline{2}$) was the impact probability (conditional probability) that the inherited antithrombin deficiency occurred under the condition of increased distribution probability, and 52 mutations had such an effect. (ii) $P(\overline{1} | \overline{2})$ was the impact probability that no disease was documented under the condition of increased distribution probability, and 3 mutations worked in such a manner. (iii) P(1|2) was the impact probability that the inherited antithrombin **I** deficiency occurred under the condition of decreased distribution probability, and 33 mutations played such a role. (iv) $P(\overline{1}|2)$ was the impact probability that no disease was documented under the condition of decreased distribution probability, and 3 mutations fell into this category. At the level of combined events, the combined results with their frequency could be seen.

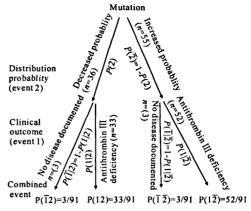


Fig. 1 Cross-impact relationship among mutations, clinical outcome, and their combination

Table 2 listed the calculated probabilities in the scheme in Fig. 1, from which several interesting points were found. (i) As $P(\overline{2})$ was larger than P(2), a mutation had six tenths chance of increasing the distribution probability in mutant antithrombin **I**. (ii) As $P(1|\overline{2})$ was 17 times of $P(\overline{1}|\overline{2})$, a mutation that increased the distribution probability had 0.95 chance of causing the inherited antithrombin **II** deficiency. (iii) As P(1|2) was remarkably larger than $P(\overline{1}|2)$, a mutation that decreased the distribution probability had more than nine tenths chance of causing the inherited antithrombin **II** deficiency.

Table 2 Probability in cross-impact analysis in Fig. 1

Probability	Calculation from the data in Fig. 1	Result
P(2)	36/91	0.3956
$P(\overline{2})$	1 - P(2) = 1 - 0.3956 or $55/91$	0.6044
$P(1 \overline{2})$	52/55	0.9455
$P(\overline{1} \mid \overline{2})$	$1 - P(1 \mid \overline{2}) = 1 - 0.9455 = 3/55$	0.0545
P (1 2)	33/36	0.9167
$P(\overline{1} 2)$	1 - P(1 2) = 1 - 0.9167 or $3/36$	0.0833
$P(1\overline{2})$	$P(1 \overline{2}) \times P(\overline{2}) = 52/55 \times 55/91 = 52/91$	0.5714
$P(\overline{1} \ \overline{2})$	$P(\overline{1} \mid \overline{2}) \times P(\overline{2}) = 3/55 \times 55/91 = 3/91$	0.0330
P(12)	$P(1 2) \times P(2) = 33/36 \times 36/91 = 33/91$	0.3626
$P(\overline{1}2)$	$P(\overline{1} 2) \times P(2) = 3/36 \times 36/91 = 3/91$	0.0330

To this point, the Bayes' $law^{[29]}$, $P(1|2) = P(2|1) \frac{P(1)}{P(2)}$, could be applied to determine the probability, P(1) that was the occurrence of inherited antithrombin I deficiency under a mutation. For this equation, P(2) and P(1|2) could be found in cross-impact analysis, while P(2|1) was the probability that the distribution probability

decreased under the condition of being inherited antithrombin II deficiency. Thus, we had P(1|2) =33/36=0.9167 (Table 2), and P(2|1)=33/(52+33) = 0.3882, so $P(1) = \frac{P(1|2)}{P(2|1)}P(2) =$ $\frac{0.9167 \times 0.3956}{0.3882} = 0.9342$, namely, a person had a chance larger than 90% of being inherited antithrombin II deficiency when a mutation was found in human antithrombin II. At present, the P(1) calculated by our approach was very meaningful for clinical settings and benefit for diagnosis as well as counseling^[30,31].

3 Conclusions

We demonstrated how to use the amino-acid distribution probability to convert a protein sequence into a numeric datum and provided a descriptively probabilistic method to couple the genotypephenotype relationship of single gene disorder. The results showed that the probability of being inherited antithrombin II deficiency was larger than 90% under a antithrombin II mutation, which would benefit the clinical diagnosis.

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